Procuring Commodities: Request for Quote or Reverse Auctions?¹

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Abstract: We examine the relative performances of reverse auctions and request for quotes in a simple commodity environment. Enterprises embarking on a reverse auction initiative often start with their commodity purchases. We conduct laboratory experiments and find that this is a poor starting point. Both the mean and variance of prices when sourcing through reverse auctions. With respect to the general investigation of auctions, the request for quote is the mirror image of a first price sealed bid auction and has the same symmetric Nash equilibrium. However, the request for quote allows identification of simple behavioral rules such as always bidding a percentage of your signal, which is indistinct from Nash equilibrium strategies in the sell auction counterpart. Consequently we estimate that one-fourth of the subjects follow a simple mark-up rule and approximately two-thirds follow a strategic Nash equilibrium strategy.

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I. Introduction

Since the mid-nineties a growing number of procurement organizations are adopting reverse auctions in efforts to reduce costs and improve performance. The competitive nature of the reverse auction is highly attractive with potential suppliers successively reducing prices until margins are squeezed to zero and then exiting the auction. This would appear to be a significant innovation over a traditional request for quote (RFQ hereafter) in which all potential suppliers submit prices incorporating positive profit margins. In line with this intuition, many early adopters of reverse auctions reported significant savings of twenty percent or greater.2

When initiating the use of reverse auctions the first question procurement organizations asks is, “where to start?” Since procurement organizations are usually responsible for sourcing a variety of goods and services, it is likely neither feasible nor prudent to start using reverse auctions for all purchases. Given the belief that reverse auctions drive prices lower, the conventional wisdom advocated in industry, for example Purchasing.com (2001) and Chafkin (2007), and academic literatures, Jap (2002) and Elmaghraby (2007), is that a new adopter should start with commodities - those purchases for which price is the sole differentiating attribute between potential suppliers. We demonstrate the opposite is true: commodities are likely a poor choice for starting a reverse auction initiative.

In this study we provide an empirical evaluation of the relative performance of reverse auctions versus RFQ’s in the procurement of commodity goods. We start by formulating hypotheses from a review of the theoretical and experimental literatures on auctions. A brief summary of these hypotheses is that the expected purchase price will be the same or lower in the RFQ than the reverse auction (depending upon the suppliers’ risk attitudes), and that purchase price variance will also be lower in the RFQ.

The development of our hypotheses involves indentifying that the RFQ is the mirror image of a first price sealed bid sell auction, and that the Nash equilibrium of the two decision problems is the same with a simple change of variable. But this change of variable, leads to a clearer separation of behavioral rules than found in previous studies of first price sealed bid auctions. These rules are bidding with a constant mark-up and bidding with according to a Nash equilibrium strategy.

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2 Two academic case studies reporting forty-three and twenty percent savings respectively are Kinney (2000) and Stein (2003.)
We test our hypotheses in controlled laboratory experiments. It would be difficult to evaluate our hypotheses with field data as different transactions may be for different commodities, suppliers’ costs are unknown, the number of potential suppliers may also be changing, and bidding and price data may be incomplete. Our laboratory experiment allows for an unconfounded evaluation of the differences in performance that solely result from changing to a reverse auction from an RFQ.

Our experimental results show that the average purchase price is significantly lower in the RFQ, suggesting commodities are not the place to start a reverse auction initiative. But why are there so many cases reporting significant savings? While the RFQ generates a lower average price, the prices also have a lower variance. Therefore we are more likely to see the reverse auction, and its more volatile prices, generate big price savings – and big losses. And perhaps there is a bias for promoting those cases of large savings.

In terms of our hypotheses, our results are most consistent with the performance predictions given by Nash equilibrium bidding models which incorporate risk-averse suppliers. However, a more thorough inspection of individual bidding behavior shows that while two-thirds of the suppliers bid consistent with a Nash bidding strategy, another quarter of the suppliers either demand a constant proportional or absolute margin. Previous studies on first price sealed bid auctions could not identify this heterogeneity.

As significant our result that commodity buys may not be suited for reverse auction purchase is for procurement officials, we believe that the demonstration of the relevance and effective use of experimental investigations and game theoretic analysis is even more valuable. To this end, in our concluding section, we provide a discussion on how our results may differ as we change the underlying aspects of our procurement environment to reflect different procurement scenarios through the use of existing current theory and experimental evidence.

II. Analytical Analysis and Development of Hypotheses

Let’s start by defining a simple commodity procurement situation. A procurement official’s task is to purchase an indivisible unit of a commodity as cheaply as possible. There are $n$ potential suppliers indexed by $i$. Each supplier can provide a unit of the commodity for the cost of $c_i$, which is incurred only if they supply the unit. The cost $c_i$ is only known by supplier $i$ – i.e. it is private information, and will typically vary across suppliers. Suppliers are symmetric in
that none has an ex ante cost advantage. Specifically, each of the supplier’s costs are drawn independently from a uniform distribution on the interval \([c_L, c_H]\). A supplier will know his own realized costs, but only the distribution of the other suppliers’ costs. The procurement official only knows that each supplier’s cost is drawn independently from this uniform distribution.

We consider two sourcing methods for this scenario. In one formulation of a reverse auction an initial high price is selected and all \(n\) suppliers are in the auction. Then the auctioneer gradually reduces the price. At any point a supplier can exit the auction, but that decision is irreversible. The number of remaining suppliers and current price are always publicly posted. When the second to last supplier exits, leaving a sole remaining supplier, the auction closes. The last remaining supplier wins the auction and receives the closing auction price, \(p\). The winning seller receives a profit of \(p - c_i\), and all other suppliers receive zero profit.\(^3\)

Game theoretic analysis reveals that the supplier has a weakly dominant strategy: remain in the auction as long as the price is greater than the supplier’s unit cost and exit when the price equals cost.\(^4\) The obvious intuition here is that all losing suppliers are driven to the point where their profit margin is zero, and revealing their true costs, before exiting the auction. Also, the supplier with the lowest realized cost wins the auction and receives a price equal to the second lowest realized cost.

In a RFQ, each supplier privately submits a price. After collecting these prices, the procurement official purchases from the lowest priced supplier at his submitted terms. In the case of a commodity, the RFQ is the procurement equivalent to a first price sealed bid auction to sell an indivisible unit.\(^5\) The pure strategy symmetric Nash Equilibrium of this RFQ calls for a supplier to submit a price according to the following function of realized costs and the number of potential suppliers (Vickrey 1961):

\[
p_l(c_i) = \frac{c_H + (n-1)c_i}{n}.
\]

\(^3\) There are other formats of the reverse auction. For example, in an open outcry format individual suppliers can announce successively lower bids until there is no supplier willing to improve upon the current existing price. The supplier submitting the last price wins the auction at that price. In the strategic analysis, bidder behavior is the open outcry and version we describe.

\(^4\) One can find standard arguments for this weakly dominant strategy in texts like (Krishner, 2002). A weakly dominant strategy in this context means that regardless of the other suppliers’ strategy there is never an instance in which the supplier can strictly increase his expected payoff by deviating from this strategy.

\(^5\) To see the equivalence, define the supplier’s private value as \(c_H - c_i\), the potential amount of value he can offer below the highest possible cost, and define the bid as, \(c_H - p\), the amount of value the supplier actually offers to the buyer.
This pricing strategy has an interesting behavioral interpretation; a supplier’s price is the expected second lowest realized cost conditional upon his cost being the lowest. Now consider the reverse auction. The winning supplier, who has the lowest realized cost, receives the price equal to the second lowest realized cost. This is an example of the celebrated revenue equivalence theorem in auction theory, and forms our first hypothesis.

**Hypothesis I:** The expected prices in the RFQ and the reverse auction are the same.

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**Figure 1: Nash Equilibrium Bidding Behavior**

To get a better insight on the revenue equivalence principle, let’s consider an example with three suppliers whose costs are independent and uniformly distributed on the interval [0, 20]. In this case the expected value of the lowest, second lowest, and highest cost are five, ten, and fifteen respectively. Figure 1 depicts the expected outcome under the Nash equilibrium. In the reverse auction, the winner’s expected cost is five. Furthermore, the expected second lowest cost, and the corresponding auction price, is ten. In the RFQ, we expect the winning supplier’s cost is also five and for him to quote a price of ten, the expected second lowest cost conditional upon five being the lowest.

Of course while the expected - or average - prices are the same, the distribution of prices is not. In the reverse auction it is easy to see that that actual winning price can occur anywhere on the interval [0, 20]. On the other hand, the support of possible prices is smaller in the RFQ. If a supplier’s realized cost is zero then conditional expectation of the second lowest cost is six and two-thirds, and this is the supplier’s RFQ bid. At the other extreme with a realized cost of twenty, the conditional expectation of the second lowest cost, and the corresponding quoted price, is

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6 The revenue equivalence theorem was first proven by Vickery (1961) and is applicable to our scenario. It was proven for a wider class of scenarios by Myerson (1981). For our concerns the version we use here states that if sellers are risk neutral, have independent and symmetric costs, and payment is function of the bid only, then the RFQ and the reverse auction will have the same expected price.
twenty. Clearly the distribution of prices in the reverse auction is a mean preserving spread of the prices in the RFQ. In fact, Vickery (1962) shows that with independent and uniformly distributed costs the variance of the price in the reverse auction is \( \frac{2(n-1)}{(n+2)(n+1)^2} (c_H - c_L)^2 \) and in the RFQ is \( \frac{(n-1)^2}{n(n+2)(n+1)^2} (c_H - c_L)^2 \). So the variance in the reverse auction price is greater than that of the RFQ by a factor of \( 2n/n - 1 \). This is our second hypothesis:

**Hypothesis II**: The variance of reverse auctions prices is greater than that of the RFQ. In this example – which will be the basis of our experiment – the variance is tripled.

With respect to the suppliers, the expected profit will be the same in both the reverse auction and the RFQ; the expected lowest cost less the expected second lowest cost multiplied by the probability of being the lowest cost supplier. In our example, this is $1.67. However, the difference is the variance of expected payoff is much higher than for the procurement official.

The variances for the supplier profit in the reverse auction is \( \frac{n}{(n+2)(n+1)^2} (c_H - c_L)^2 \) and in the RFQ is \( \frac{1}{n(n+2)(n+1)^2} (c_H - c_L)^2 \); the difference is a factor of \( n^2 \). The volatility of the supplier’s profit is much greater than the volatility of the buyer’s benefit. We summarize these observations with the following pair of hypotheses.

**Hypothesis Ia**: The expected profit of a supplier is the same in the RFQ and the reverse auction.

**Hypothesis Iia**: The variance of a supplier’s profit in a reverse auctions prices is greater than that of the RFQ. In this example, the variance is nine times greater.

We now formulate our next hypothesis by relaxing the constraint that all suppliers are risk neutral. Clearly when entering into a procurement process the outcomes are uncertain to the suppliers. However, in the reverse auction the presence of risk averse suppliers predicted behavior or the resulting prices do not change, because of the weakly dominant strategy. This is not true for the RFQ. Holt (1980) shows that if all suppliers have the same risk averse von Neumann-Morgenstern expected utility function, then there exists a symmetric Nash equilibrium in which the expected price is lower than that of the reverse auction.

This issue has been quite prominent in the experimental economics literature on single unit private value sealed bid sell auctions. The experiments addressed in that literature are a mirror image situation to the one we consider. An auctioneer wishes to sell a single indivisible
unit of a good and there are \( n \) possible buyers. Each buyer draws an independent private valuation from a uniform distribution on the interval \([0, v_H]\). This valuation is an individual specific price that he can resell the object to the experimenter for. The symmetric Nash equilibrium bidding strategy for this auction is

\[
B(v_i) = \frac{n-1}{n} v_i.
\]

If we set the number of buyers to three and the highest possible valuation to twenty, we have the selling counterpart to our example. The equilibrium bidding function is depicted in Figure 2.

![First Price Sealed Bid Auction](image)

Figure 2: Nash Equilibrium Bidding Behavior First Price Sealed Bid Auction

Early studies, such as Coppinger, Smith, and Titus (1980) and Cox, Roberson, and Smith (1982), found that the vast majority of bids lay above the Nash equilibrium bidding line in Figure 2 and below the forty-five degree line at which bid equals value – in the area labeled the cone of risk averse bids. Consequently, they found that the first price sealed bid auction provided the buyer with higher average prices than those received in the English outcry auction (the selling counterpart to the reverse procurement auction.) Cox, Roberson, and Smith suggest that these choices are the result of heterogeneous risk attitudes amongst the bidders.\(^7\) Further they propose a Nash equilibrium model in which a bidder’s risk attitude is characterized by a single parameter that is private information.\(^8\) Given Holt’s theoretical model of the RFQ with risk averse suppliers and the strong empirical evidence from experiments of first price sealed bid sellers auctions we offer the following hypothesis

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\(^7\) Other alternative models to explain this observation have been recently proposed such as regret theory, Filiz-Ozbay (2007), and directional learning theory, Neugebauer (2006).

\(^8\) There have been numerous studies which have addressed the appropriateness of the risk aversion explanation; Kagel and Levin (2008) provide a survey of these criticisms.
**Hypothesis III:** The expected or average price in the RFQ will be lower than in the reverse auction.

Our RFQ investigation provides a unique opportunity, although tangential to our practical concern, to conduct a powerful test amongst competing behavioral hypotheses in the private value first price sealed bid auction. Notice in Figure 2 that the Nash equilibrium bid function also corresponds to the behavioral rule always bid a fixed proportion of your realized value. If we allow for heterogeneous risk attitudes then any estimated linear bid function, which goes through the origin and has slope less than one, is consistent with both fixed proportion of realized value and a Nash Equilibrium bid functions.\(^9\)

Now consider the RFQ Nash equilibrium bid function depicted in Figure 1. In this case, these two behavioral rules generate quite distinct behavior. A Nash equilibrium bidder, demands zero margin when his has the worst (highest) cost realization, and his margin demanded, both proportional and absolute, increases as realized cost decreases. This pricing behavior is strikingly different from that of a supplier who simply demands a fixed percentage margin, i.e. their pricing strategy is simply \(p_t(c_i) = (1 + x)c_i\). In our data analysis, we also consider a third behavioral rule in which a supplier demands a constant absolute margin, or profit level. For this type of supplier his pricing function is of the form \(p_t(c_i) = \alpha + c_i\). To summarize we have three hypothesis regarding supplier behavior in the RFQ which generate distinct choices.

**RFQ Bidding Hypotheses:** A supplier follows either a i) Nash Equilibrium strategy, ii) a constant proportional margin demanding strategy, or iii) a constant absolute margin demanding strategy.

### III. Experimental Design and Procedures

We ran our experimental sessions at the Economics Laboratory at UCSD and at the NUS School of Business. All participants were either undergraduate or master level students at one of the two universities. RFQ and reverse auctions were conducted at both facilities. The number of participants in an experimental session consisted of some multiple of three between nine and eighteen. Experimental sessions lasted no more than one hour and subject earned between $8 (S$12) and $64 (S$96.)

\(^9\) Note that in the Holt model all bidders have the same risk attitude and would have the same bid function. Meanwhile, in the Cox, Roberson and Smith model subjects have heterogeneous risk attitudes and will have varying linear bidding strategies with one caveat. Above the threshold of the maximum possible bid of a risk neutral buyer, the bid function is strictly concave.
We adopted the simple previous three supplier example as the basis for our experiments. In a session, the subjects participated in series of thirty-two rounds of either RFQ’s or reverse auctions. In each round, the subjects were randomly re-partitioned into a set of triads. The first two rounds were for practice; the subjects earned no money and we do not report the data. For the remaining rounds the participants earnings were given in an experimental currency. The exchange rates were one experimental dollar to $0.33 (or S$0.50.) In total we conducted on 720 RFQ’s and 360 reverse auctions.10

All participants’ decisions in the experiment were made through a personal computer running a custom designed software program, and in a partition designed to ensure private decisions. In the RFQ, at the start of each period the subject is revealed his or her realized private cost (a new cost was drawn each period.) The subject was also shown the period number and the number of other subjects in the RFQ. The subject was prompted to submit a price (restricted to be between zero and thirty), but could take as much time desired to do so. After all prices were submitted, the auction results were revealed. These results consisted of the bid submitted by the subject, the amount of the winning bid, and the subject’s period profit. All of the information was then entered into a display, along with the cumulative profit, at the bottom of the computer screen for future reference. At the conclusion of the experiment, subjects were informed of their total earnings and paid that amount privately.

The reverse auction experiment had the same procedures, except for the execution of auction. On the computer screen there was a display of the current auction price and a button that one could press to exit the auction. The auction started with an initial price of $21. Then the price was decremented at the rate of ten cents every half a second. As auction participants exited the auction, subjects could observe the decrement of the displayed number of participants remaining in the auction. At the close of the auction their cost, exit price, period profits, and cumulative profits were entered in history viewing area.

IV. Data Analysis and Results

We first address the relative performance of the two alternative procurement procedures, and observe that procurement costs are lower and less volatile in the RFQ. For suppliers, average profits are higher, but at the same time, more volatile in the reverse auction.

10 Due to the strong agreement of the data with the weakly dominant bidding strategy in the reverse auction we collected an unbalanced sample.
Figure 3 presents a histogram of prices in the two mechanisms: realized prices are collected into two-dollar wide bins and we report the percentage for each bin. Notice that the two distributions are quite distinct. The RFQ price histogram has a single peak, a mode below ten, and has a relatively smooth, although skewed, distribution. While the histogram of reverse auction prices exhibits more dispersion and irregularity. Clearly the price distributions are not the same.

Table 1: Summary Statistics on Prices and Profits

<table>
<thead>
<tr>
<th></th>
<th>Mean Price</th>
<th>Price Variance</th>
<th>Mean Supplier Profit</th>
<th>Supplier Profit Variance</th>
</tr>
</thead>
<tbody>
<tr>
<td>Reverse Auction</td>
<td>10.60</td>
<td>19.11</td>
<td>1.68</td>
<td>10.67</td>
</tr>
<tr>
<td>RFQ</td>
<td>7.90</td>
<td>9.63</td>
<td>0.90</td>
<td>3.04</td>
</tr>
<tr>
<td>Difference</td>
<td>-25.5%</td>
<td>-50.4%</td>
<td>-46.3%</td>
<td>-71.5%</td>
</tr>
</tbody>
</table>

Inspecting the mean and variance of the prices reveals that RFQ delivers lower cost and lower price volatility to the procurement official. Table 1 present the means and variances of the two price distributions and the differences in these values. The first column shows the mean, the standard error of the means, and the percentage difference in prices. Two-tailed $t$-tests lead us to reject that the reverse auction mean price (5% level of significance) and the RFQ mean price (1% level of significance) are equal to the risk neutral Nash equilibrium prediction of ten. Moreover,
we reject that the mean prices are the same in favor of the hypothesis that the RFQ mean price is lower by conducting a $t$-test for unequal variances (1% level of significance.)

**Result 1:** Procurement prices are lower, 25.5% in fact, for the RFQ. We reject hypothesis I in favor of hypothesis III.

In the second column we report information for the variance on prices. Under the risk neutral Nash equilibrium bidding models, the variance of the prices in the reverse auction and RFQ are twenty and six and two-thirds respectively. We give the sample variances and beneath those we give bootstraps of the 95% confidence intervals of the variance under Nash equilibrium. We can’t reject that the variance is equal to the predicted value in the reverse auction, but we do reject this for the RFQ as the estimated variance is well outside and above the confidence interval. When comparing the variances, we reject that hypothesis they are the same in favor of the hypothesis that the reverse auction price have greater variance (Levin $F$-tests for homogeneous variances at the 1% level of significance.) However, the price variance in the reverse auction is only twice as large (i.e. the -50.4% change noted in the table,) not triple as predicted in theory.

**Result 2:** Procurement costs are more volatile in the reverse auction. Our experimental results are consistent with hypothesis II. Thus, unless a procurement organization is very risk loving, the lower average price and lower price variability make the RFQ the better method in the commodity situation.

What about the welfare of the suppliers? As we see in the third column, we can’t reject ($t$-test at 1% level of significance) that the mean supplier profit in the Reverse auction is equal to the theoretical value of 1.67. Average supplier profit in the RFQ is 46.3% lower, and we reject these profits are the same as the risk neutral Nash equilibrium bidding and the empirical mean of the average profit in the reverse auction experiments (both at the 1% level of significance.) While suppliers actually do better on average in the reverse auctions, the volatility of the payoffs are about two and a half times more volatile.

**Result 3:** Supplier’s profits are higher in the reverse auction rather than the RFQ and we reject hypothesis Ia. And we also find these profits are more volatile in the reverse auction as well – confirming hypothesis Iia.

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11 Since the distribution of prices under Nash equilibrium in know to us and non-normal simple $t$-tests of the estimated variances are not valid. So, we bootstrap with 10,000 draws from the known distribution and then calculate the 95% confidence intervals for the theoretical values of the variance.
The relative performances of the two procurements methods are most consistent with risk averse Nash Equilibrium models. These results show promise that the robustness of this type of model would extend to other type of procurement and supply chain activities. But this model is formulated as a description of individual level behavior. So, this robustness likely only holds if the model explains what happens at the individual level.

Figure 4: Reverse Auction Exit Price Versus Costs

With respect to the reverse auction, risk aversion does not affect the weakly dominant strategy and we should observe those who exit an auction do so at their true costs. Figure 4 plots seller’s exit prices versus there realized costs. While much of the data does adhere closely to forty-five degree line as we would expect, there is a surprising amount that doesn’t. Much of these occurrences are at high cost levels. There are a number of observations were a seller opts out as soon as possible when the auction opens at the price of $21; perhaps frustrated at receiving a high cost level with little opportunity to earn money. However, there is another story for the number of exit prices below cost. In most of these cases, an individual remained in the auction while there are still two other suppliers remain. Then, as soon as another supplier exited, this individual exited as well. To see the effect of how binding behavior depends on cost, we plot the
realized auction price versus the second lowest realized costs in Figure 5. Here we can see much crisper conformity with the forty-five degree line. To quantify this we present an OLS fitted trend line through the origin. The slope coefficient is essentially one and this regression explains over 93% of the variation as indicated by the $R^2$ statistic.

![Figure 5: Reverse Auction Price Versus Second Lowest Costs](image)

Now we turn our attention to the individual subject behavior in the RFQ procurement sessions. The data in these sessions, when pooled, are consistent with the story of a Nash Equilibrium bidding model with heterogeneity in risk attitudes, and some aspects of previous experiments on first price sealed bid auctions. However, we show that the procurement framing of the private value auction allows us to identify some bidding rule heterogeneity. One-fourth of the subjects’ quotes reflect constant or absolute margin demands rather than strategic margin demands found in Nash equilibrium bidding functions – namely for lower cost (and a more competitive position) one demands a higher margin.
Figure 6 displays the 2160 submitted RFQ prices versus realized cost. We also have plotted the risk neutral Nash Equilibrium pricing function, the price equal to cost line (or the forty-five degree line), and an OLS fitted line for the data. Clearly, the majority of bids are above cost but below the risk neutral Nash equilibrium price - just as a risk averse Nash model predicts. However, there are also many observations outside (particularly above) this cone of risk averse actions. This begs the question if there is some type of bidder heterogeneity beyond risk attitudes generating this behavior.

We explore to what extent competing behavior models best explain individuals’ pricing behavior, and we find that one-fourth of the subjects use a simple mark-up rule as opposed to a more strategic rule or other linear rule. We analyze this heterogeneity by first considering a linear pricing rule

(1) \[ p_t = \alpha + \beta \cdot \text{cost}_t + \varepsilon_t, \]

where \( p_t \) is the RFQ price submitted in round \( t \), \( \text{cost}_t \) is the suppliers realized cost in round \( t \), and \( \varepsilon_t \) is a random unbiased shock. We estimate this model for each of our seventy-two subjects by OLS.
Two types of non-strategic behavioral models are a special case of (1). If a supplier in our experiment follows a rule of always demanding a constant percentage margin, then his pricing rule is

\[ p_t = \beta \cdot \text{cost}_t + \varepsilon_t, \]

This is a special case of (1) with \( \alpha = 0 \), \( \beta = (1 + x) \) where \( x \) is the percentage mark-up over cost. A simple hypothesis test on the significance of \( \alpha \) in the estimation of (1) allows us to reject the constant percentage mark-up rule in favor of the general linear model. Now if a subject always submits a price that includes a constant absolute profit level, the his pricing rule is

\[ p_t = \alpha + \text{cost}_t + \varepsilon_t, \]

One way to test for this model is to conduct a hypothesis test for \( \beta = 1 \) in (1).

If a supplier prices according to a risk averse Nash equilibrium what are the restrictions\(^{12}\) on the pricing function? The supplier bids linearly, also demands zero margin at the highest possible cost, $20, and the price function will lie in the cone defined by the risk neutral bidding function and price at cost. Mathematical, these are the following restriction on (1): \( \alpha + \beta \cdot 20 = 20 \), and \( \beta \epsilon[n - 1/n, 1] \). We estimate this model with the linear condition via restricted OLS, but we don’t constraint the slope coefficient. Rather we observe if the estimated value lies in the appropriate range. One can test this model against the general linear pricing model with an \( F \)-test comparing the two models’ sum-squared residuals.

We develop a classification for the subjects in our experiment by estimating the four possible models for each subject, and then performing the appropriate hypothesis test for each specific pricing model versus the general linear one. This procedure leaves the possibility of a subject being selected for more than one model; however, that was not the case as we either rejected all three alternative models, or failed to reject for exactly one.

The results show that out of seventy-two subjects; forty-five follow Nash equilibrium pricing, fourteen choose prices with constant absolute margins, three set prices with a constant proportional margin, and eight are best described by an unrestricted linear pricing rule. The

\(^{12}\) If we assume that Nash equilibrium requires a common belief over the distribution of risk attitudes as well as the distribution over costs, then we would only observe the linear pricing function for the setting of homogeneous constant relative risk aversion (CRRA) and then all pricing functions would be the same. Cox, Smith and Roberson [1982] develop a model allowing for heterogeneous CRRA which generates distinct pricing functions that are linear above some price threshold. However, our approach is more appropriate if we don’t assume a common belief over the distribution of risk attitudes allowing for a broader equilibrium set of strategies. Ledyard [1986] makes this point for an extensive class of strategic situations.
details of these results are reported in Tables 2-5. In these tables we report the estimated intercept and slope coefficients – with the standard errors when appropriate, the predicted price when cost is twenty, and the $R^2$ statistic of the regression. We also present the individual estimated equations from the highest to lowest $R^2$.

In Table 2 we present the estimated pricing equations for the forty-four Nash equilibrium pricing rule adopters. First note that except for worst fitting estimation all of the slope coefficients are in the interval $[2/3, 1]$ as postulated, and correspondingly to the imposed linear restriction, all of the estimated intercept terms are less than the risk neutral value.

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<th>$\beta$</th>
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<td>20</td>
<td>0.917</td>
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<td>(.251)</td>
<td>0.838</td>
<td>-</td>
<td>20</td>
<td>0.973</td>
<td>4.65</td>
<td>(.467)</td>
<td>0.77</td>
<td>-</td>
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<td>(.300)</td>
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<td>-</td>
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<td>3.44</td>
<td>(.464)</td>
<td>0.83</td>
<td>-</td>
<td>20</td>
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<td>0.818</td>
<td>-</td>
<td>20</td>
<td>0.971</td>
<td>4.18</td>
<td>(.472)</td>
<td>0.79</td>
<td>-</td>
<td>20</td>
<td>0.901</td>
</tr>
<tr>
<td>2.309</td>
<td>(.266)</td>
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<td>-</td>
<td>20</td>
<td>0.967</td>
<td>4.65</td>
<td>(.460)</td>
<td>0.77</td>
<td>-</td>
<td>20</td>
<td>0.900</td>
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<td>0.853</td>
<td>-</td>
<td>20</td>
<td>0.966</td>
<td>4.67</td>
<td>(.548)</td>
<td>0.77</td>
<td>-</td>
<td>20</td>
<td>0.900</td>
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<td>5.523</td>
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<td>0.724</td>
<td>-</td>
<td>20</td>
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<td>4.60</td>
<td>(.518)</td>
<td>0.77</td>
<td>-</td>
<td>20</td>
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<td>0.877</td>
<td>-</td>
<td>20</td>
<td>0.965</td>
<td>4.30</td>
<td>(.470)</td>
<td>0.79</td>
<td>-</td>
<td>20</td>
<td>0.891</td>
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<tr>
<td>4.306</td>
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<td>0.785</td>
<td>-</td>
<td>20</td>
<td>0.963</td>
<td>5.32</td>
<td>(.481)</td>
<td>0.73</td>
<td>-</td>
<td>20</td>
<td>0.884</td>
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<tr>
<td>4.311</td>
<td>(.287)</td>
<td>0.784</td>
<td>-</td>
<td>20</td>
<td>0.962</td>
<td>5.16</td>
<td>(.477)</td>
<td>0.74</td>
<td>-</td>
<td>20</td>
<td>0.878</td>
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<tr>
<td>2.402</td>
<td>(.278)</td>
<td>0.88</td>
<td>-</td>
<td>20</td>
<td>0.961</td>
<td>4.29</td>
<td>(.437)</td>
<td>0.79</td>
<td>-</td>
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<td>4.019</td>
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<td>0.799</td>
<td>-</td>
<td>20</td>
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<td>5.48</td>
<td>(.403)</td>
<td>0.73</td>
<td>-</td>
<td>20</td>
<td>0.867</td>
</tr>
<tr>
<td>2.919</td>
<td>(.283)</td>
<td>0.854</td>
<td>-</td>
<td>20</td>
<td>0.958</td>
<td>5.54</td>
<td>(.359)</td>
<td>0.72</td>
<td>-</td>
<td>20</td>
<td>0.856</td>
</tr>
<tr>
<td>3.888</td>
<td>(.253)</td>
<td>0.806</td>
<td>-</td>
<td>20</td>
<td>0.958</td>
<td>5.57</td>
<td>(.449)</td>
<td>0.72</td>
<td>-</td>
<td>20</td>
<td>0.855</td>
</tr>
<tr>
<td>2.752</td>
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<td>0.862</td>
<td>-</td>
<td>20</td>
<td>0.957</td>
<td>6.36</td>
<td>(.814)</td>
<td>0.68</td>
<td>-</td>
<td>20</td>
<td>0.741</td>
</tr>
<tr>
<td>4.095</td>
<td>(.281)</td>
<td>0.795</td>
<td>-</td>
<td>20</td>
<td>0.952</td>
<td>4.92</td>
<td>(.920)</td>
<td>0.75</td>
<td>-</td>
<td>20</td>
<td>0.690</td>
</tr>
<tr>
<td>3.279</td>
<td>(.262)</td>
<td>0.836</td>
<td>-</td>
<td>20</td>
<td>0.95</td>
<td>9.69</td>
<td>(.818)</td>
<td>0.52</td>
<td>-</td>
<td>20</td>
<td>0.679</td>
</tr>
<tr>
<td>5.899</td>
<td>(.408)</td>
<td>0.705</td>
<td>-</td>
<td>20</td>
<td>0.937</td>
<td></td>
<td></td>
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</tr>
</tbody>
</table>
One of the advantages of our experimental design is that the procurement formulation gives distinct predictions for Nash equilibrium choices and for pricing at a constant proportion of one’s private signal. We only classify three subjects as using such a rule, and their estimated pricing rules are presented in Table 3. We can see there is substantial variation in the estimated mark-ups and the fit, as indicated by the $R^2$ statistic, is poorer than in other models. Table 4, reports the estimated mark-up for those who price with a fixed absolute profit margin. These subjects who bid with constant absolute margins generate a substantial amount of data observed outside the predicted regions. Also contributing to these observations are those subjects who we conclude are not using one of the more refined behavior models. The estimated linear pricing rules for these subjects are reported in Table 5.

Table 3: Estimation for constant proportional margins

\[ p(\text{Cost}) = \alpha + \beta \times \text{Cost} + \epsilon; \text{ where } \alpha=0 \]

<table>
<thead>
<tr>
<th>$A$</th>
<th>standard error</th>
<th>$B$</th>
<th>standard error</th>
<th>$p(20)$</th>
<th>$R^2$</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.00</td>
<td>-</td>
<td>1.13</td>
<td>(.046)</td>
<td>22.68</td>
<td>0.791</td>
</tr>
<tr>
<td>0.00</td>
<td>-</td>
<td>1.41</td>
<td>(.075)</td>
<td>28.26</td>
<td>0.593</td>
</tr>
<tr>
<td>0.00</td>
<td>-</td>
<td>1.29</td>
<td>(.076)</td>
<td>25.86</td>
<td>0.540</td>
</tr>
</tbody>
</table>
Table 4: Estimations for fixed absolute margin

\[ p(Cost) = \alpha + \beta(Cost); \text{ where } \beta = 1 \]

<table>
<thead>
<tr>
<th>( A )</th>
<th>standard error</th>
<th>( \beta ) standard error</th>
<th>( p(20) )</th>
<th>( R^2 )</th>
</tr>
</thead>
<tbody>
<tr>
<td>1.08</td>
<td>.022</td>
<td>1.00</td>
<td>21.08</td>
<td>0.998</td>
</tr>
<tr>
<td>2.23</td>
<td>.067</td>
<td>1.00</td>
<td>22.23</td>
<td>0.984</td>
</tr>
<tr>
<td>0.91</td>
<td>.096</td>
<td>1.00</td>
<td>20.91</td>
<td>0.969</td>
</tr>
<tr>
<td>1.01</td>
<td>.088</td>
<td>1.00</td>
<td>21.01</td>
<td>0.966</td>
</tr>
<tr>
<td>1.69</td>
<td>.109</td>
<td>1.00</td>
<td>21.69</td>
<td>0.961</td>
</tr>
<tr>
<td>2.59</td>
<td>.138</td>
<td>1.00</td>
<td>22.59</td>
<td>0.952</td>
</tr>
<tr>
<td>2.26</td>
<td>.143</td>
<td>1.00</td>
<td>22.26</td>
<td>0.945</td>
</tr>
<tr>
<td>4.05</td>
<td>.149</td>
<td>1.00</td>
<td>24.05</td>
<td>0.937</td>
</tr>
<tr>
<td>4.70</td>
<td>.151</td>
<td>1.00</td>
<td>24.70</td>
<td>0.928</td>
</tr>
<tr>
<td>3.18</td>
<td>.173</td>
<td>1.00</td>
<td>23.18</td>
<td>0.875</td>
</tr>
<tr>
<td>2.39</td>
<td>.230</td>
<td>1.00</td>
<td>22.39</td>
<td>0.848</td>
</tr>
<tr>
<td>3.55</td>
<td>.249</td>
<td>1.00</td>
<td>23.55</td>
<td>0.835</td>
</tr>
<tr>
<td>5.38</td>
<td>.330</td>
<td>1.00</td>
<td>25.38</td>
<td>0.702</td>
</tr>
<tr>
<td>1.23</td>
<td>.310</td>
<td>1.00</td>
<td>21.23</td>
<td>0.693</td>
</tr>
<tr>
<td>5.52</td>
<td>.349</td>
<td>1.00</td>
<td>25.52</td>
<td>0.612</td>
</tr>
</tbody>
</table>

Table 5: Estimations for general linear bidders

\[ p(Cost) = \alpha + \beta(Cost) + \varepsilon \]

<table>
<thead>
<tr>
<th>( A )</th>
<th>standard error</th>
<th>( B ) standard error</th>
<th>( p(20) )</th>
<th>( R^2 )</th>
</tr>
</thead>
<tbody>
<tr>
<td>2.00</td>
<td>.224</td>
<td>0.93</td>
<td>20.67</td>
<td>0.990</td>
</tr>
<tr>
<td>3.89</td>
<td>.269</td>
<td>0.75</td>
<td>18.93</td>
<td>0.969</td>
</tr>
<tr>
<td>5.29</td>
<td>.473</td>
<td>0.81</td>
<td>21.43</td>
<td>0.936</td>
</tr>
<tr>
<td>6.11</td>
<td>.507</td>
<td>0.81</td>
<td>22.38</td>
<td>0.928</td>
</tr>
<tr>
<td>4.37</td>
<td>.572</td>
<td>0.86</td>
<td>21.56</td>
<td>0.925</td>
</tr>
<tr>
<td>7.17</td>
<td>.369</td>
<td>0.53</td>
<td>17.83</td>
<td>0.897</td>
</tr>
<tr>
<td>4.37</td>
<td>.530</td>
<td>0.69</td>
<td>18.21</td>
<td>0.875</td>
</tr>
<tr>
<td>7.39</td>
<td>.990</td>
<td>0.28</td>
<td>13.00</td>
<td>0.290</td>
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</tbody>
</table>
V. Discussion

Our experiments provide a clear demonstration of the relative performance of the RFQ versus the reverse auction in a commodity procurement setting. In light of these results, any procurement organization should proceed cautiously if initiating a reverse auction strategy with the purchase of commodities. From the procurement organization’s perspective reverse auctions not only deliver higher expected prices, but also greater price variability. It is worth noting that increased price variability, has a negative impact beyond a procurement organization’s distaste of increased price uncertainty. Recent research studies, Jap [2007] and Haruvy and Jap [2008], have documented that relationship between suppliers and buyers is negatively impacted by the adoption of the reverse auction. The increase in price volatility we document can be at least a partial source of this seller animosity.

Beyond summarizing our results, we want to consider why our results contradict the frequently reported successes in first trials of reverse auctions? The issue of price variance is again a possible source. Higher price volatility generates a larger number of great successes when switching from the RFQ to the reverse auction, of course there will also be more spectacular failures as well. If there is a bias in reporting successful versus negative outcomes, see for example An-Wen Chan et Al (2004) and Graham, Harvey, and Rajgopal (2005), then we should see more reports of successful initial reverse auctions. However, instead of just assuming there is a reporting bias, let’s consider how our commodity sourcing experiment might differ from those in the field.

In our study we fixed the number of potential suppliers and made several strong assumptions; relaxing any of these could lead to different outcomes. In practice, it’s likely that when initiating reverse auction sourcing variables other than the auction format change. Such e-sourcing efforts typically include attempts to increase the number of qualified participating suppliers. It’s clear in both the RFQ and the reverse auction increasing the number of suppliers will reduce price, but all of our theories suggest that the RFQ will still retain its advantages.

In our analysis we assume suppliers independently draw costs from the same distribution. If we relax that assumption, for example one supplier has a clear advantage in location and likely lower cost, and then the better sourcing practice is not so evidently clear. Theoretical studies, such as Maskin and Riley (2000) and Cantillon (2008), predict that whether a reverse auction will lead to lower price depends crucially on the distributions of costs. Experiments such as Guth
et al.(2005) reveal that subjects overbid, not realizing how competitively strong their positions are at times, from theoretical bid functions in such environments. Another key element of our formulation is that costs are supplier specific and private information. However, if the true structure is such that costs are correlated or the same across firms, but each firm only has an estimate of this cost, then reverse auction provides the lower expected costs as shown in Milgrom and Roberts’ (1982) seminal paper.

One area where reverse auctions do show much promise is in the procurement of goods for which price is not the only differing attribute between suppliers. Researchers have studied two cases: when non-price attributes are exogenous and when they are determined within the auction. Engelbrech-Wiggans, Haruvy, and Katok (2007) find significant gains to the procurement official when suppliers bid on price, and then the buyer chooses the winner versus awarding the contract to the lowest bidder. The performance of the buyer determined winner auction versus the RFQ in these setting is studied in Haruvy and Katok (2008), and they find the RFQ is better if suppliers have accurate and precise information regarding the quality of other sellers. On the other hand, Shachat and Swarthout (2003), find that a reverse auction with buyer assigned bidding credits can provide better outcomes than the RFQ for both suppliers and sellers. There is also a large and promising literature on successful reverse auction examples where the quality is determined within the reverse auction, for example Chen-Ritzo et al (2005) and Parkes and Kalagnanam (2005).

In summary, the knowledge gained in our study can inform procurement organizations that commodity buys may not be the optimal place to start a reverse auction initiative. Also, for the literature on sealed bid auctions our procurement framing permitted a new identification of bidding heterogeneity across individuals. Through the use of game theoretic arguments and controlled laboratory experiments were able to justify our results. We hope that practitioners see that the value of the tandem use of these tools is for addressing other procurement problems.

References


